

7.3) Generativni modeli

75

• u log. regresiji direktno modelinemo $p_k(x)$, $k \in S$
→ primjer "diskriminacijske" metode

• alternativno, ako modelinemo $f_k(x) := P(X=x | Y=k)$,
 $x \in \mathbb{R}^p$, $k \in S$, te $\pi_k := P(Y=k)$,

$$p_k(x) = \frac{f_k(x) \pi_k}{\sum_{i \in S} f_i(x) \pi_i} \quad (7.10)$$

Bayes

• Zapravo modelinemo $P(X=x, Y=k) = f_k(x) \pi_k$, tj:
 $P((X, Y) \in \cdot) \rightarrow$ "generativne" metode

• (1) Linearna diskriminacijska analiza (LDA)

(2) Koadnutna " " " " (QDA)

(3) Naivni Bayes

Uočimo, (7.10) povlači (uz 0-1 gubitak)

$$f^*(x) = \operatorname{argmax}_{k \in S} f_k(x) \pi_k. \quad (7.11)$$

7.3.1 LDA

• Pretpostavka je da $X | Y=k \sim N(\mu_k, \Sigma_k)$, tj:

$$f_k(x) = \frac{1}{(2\pi)^{\frac{p}{2}} \sqrt{\det(\Sigma_k)}} \cdot \exp\left\{-\frac{1}{2}(x-\mu_k)^T \Sigma_k^{-1}(x-\mu_k)\right\}, x \in \mathbb{R}^p, \quad (7.12)$$

te dodatno

$$\Sigma_k = \Sigma, \forall k \in S! \quad (7.13)$$

$f^*(x) = \underset{h \in S}{\operatorname{argmax}} \log(\pi_h(x) \cdot \pi_h)$
 $= \log(\pi_h(x)) + \log(\pi_h)$

(Mahalevskijeva udaljenost od x i μ_h)

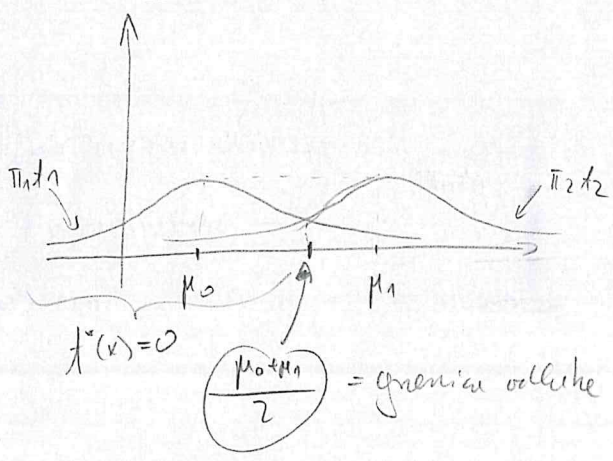
$\stackrel{(7.11)}{=} \underset{h \in S}{\operatorname{argmin}} \left\{ \frac{1}{2} (x - \mu_h)^T \Sigma^{-1} (x - \mu_h) - \log(\pi_h) \right\}, x \in \mathbb{R}^p$
 $=: d_h(x)$

Ako je $\Sigma = I_p$ te $\pi_0 = \dots = \pi_{k-1}$,

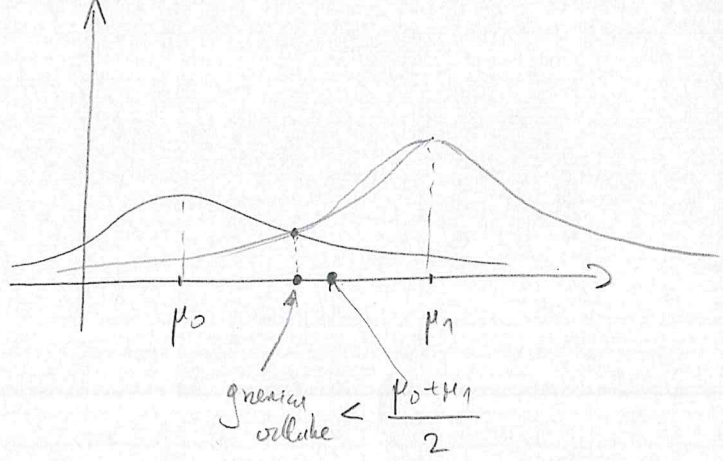
$f^*(x) = \underset{h \in S}{\operatorname{argmin}} \|x - \mu_h\|_2^2$

Pr. / $p=1, k=2$ ($\Sigma = \sigma^2 > 0$)

1^o $\pi_0 = \pi_1$



2^o $\pi_1 > \pi_2$



[uzimamo u obzir da je vise (x_i, y_i) s $|y_i = 1|$]

$d_h(x) = \frac{1}{2} x^T \Sigma^{-1} x - x^T \underbrace{\Sigma^{-1} \mu_h}_{=: w_h} + \underbrace{\mu_h^T \Sigma^{-1} \mu_h - \log(\pi_h)}_{=: -b_h}$

$f^*(x) = \underset{h \in S}{\operatorname{argmax}} x^T w_h + b_h$

linearna diskriminacijska \rightarrow "LDA"
 f je

Specijalno, LDA \rightarrow linearna granica odluke.

Parametre projekcije su:

$$\hat{\pi}_h = \frac{1}{n} \sum_{i=1}^n \mathbb{1}_{\{y_i = h\}}$$

$$\hat{\mu}_h = \frac{\sum_{i=1}^n x^{(i)} \mathbb{1}_{\{y_i = h\}}}{\sum_{i=1}^n \mathbb{1}_{\{y_i = h\}}}$$

"pooled" varijanca kov. matrica

$$\hat{\Sigma} = \frac{1}{n - (K)} \sum_{i=1}^n (x^{(i)} - \hat{\mu}_{y_i})(x^{(i)} - \hat{\mu}_{y_i})^T$$

$$\hat{f}(x) = \operatorname{argmin}_{h \in S} \hat{d}_h(x)$$

(K-1)-dim. afin prostor

Prop. 1 (i) (ii) Pretp. da je $\hat{\Sigma} = I_p$, te $H_{K-1} = \mu_0 \oplus \operatorname{span}\{\mu_1, \dots, \mu_{K-1}\}$.

$$\hat{f}(x) = \operatorname{argmin}_{h \in S} \left\{ \frac{1}{2} \|x - \mu_h\|_2^2 - \log(\hat{\pi}_h) \right\}$$

$$= \operatorname{argmin}_{h \in S} \left\{ \frac{1}{2} \|\operatorname{Proj}_{H_{K-1}}(x) - \mu_h\|_2^2 - \log(\hat{\pi}_h) \right\}$$

\Rightarrow dim je $K-1 < p$, imamo manje dimenzije

(ii) \Rightarrow općenito, prv neprenimo transformaciju

$$x \mapsto x^* := \hat{\Sigma}^{-1/2} x \quad (\Rightarrow (x - \hat{\mu}_h)^T \hat{\Sigma}^{-1} (x - \hat{\mu}_h) = \|x^* - \hat{\mu}_h^*\|_2^2, \hat{\Sigma}^* = I_p)$$

te gledamo projekcije na $H_{K-1}^* = \mu_0^* \oplus \operatorname{span}\{\mu_1^*, \dots, \mu_{K-1}^*\}$.

Može se pokušati da postoje vektori $\{v_1, \dots, v_{K-1}\} \in \mathbb{R}^p$

t. d.

$$\hat{f}(x) = \operatorname{argmin}_{h \in S} \left\{ \frac{1}{2} \|V \cdot x - V \cdot \hat{\mu}_h\|_2^2 - \log(\hat{\pi}_h) \right\}$$

za $V = \begin{bmatrix} -v_1^T \\ \vdots \\ -v_{K-1}^T \end{bmatrix} \in \mathbb{R}^{(K-1) \times p}$

$v_i^T \cdot x$ je tzv. i-ta diskriminacijska koordinata

\rightarrow Wine-LDA.R

Primjerena statistika i ESL 4.3.3

→ dve klas kod LDA, ali bez pref. $\Sigma_k = \Sigma, \forall k \in S.$

$$\Rightarrow f^*(x) = \underset{h \in S}{\operatorname{argmax}} \left\{ -\frac{1}{2} (x - \mu_h)^T \Sigma_h^{-1} (x - \mu_h) - \frac{1}{2} \log(\det(\Sigma_h)) + \log \pi_h \right\}$$

$$= \underset{h \in S}{\operatorname{argmax}} \left\{ -\frac{1}{2} x^T \Sigma_h^{-1} x + \dots \right\}$$

kvadratne granice odlike.

Ipak, u odnosu na LDA, QDA procjenjuje $\hat{\Sigma}_k, k \in S$

→ dodatnih $\frac{p(p-1)}{2} \cdot (K-1)$ parametara!

kada je p velik, čak i LDA procjenjuje previše parametara

→ Σ, μ_k, π_k
 $\frac{p(p-1)}{2} \cdot K \cdot p \quad K-1$

7.3.3 Naivni Bayes (NB)

→ pref. da su $\forall k \in S, X_1, \dots, X_p$ nezavisne uz članu $Y=k, t_j.$

$$f_k(x) = \prod_{j=1}^p f_{k,j}(x_j), \quad \forall x = (x_1, \dots, x_p) \in \mathbb{R}^p \quad (7.15)$$

[nezavisnost je oblik regulencije]

Marginalne gustoće možemo procjeniti npr.

• ako $X_j \in \mathbb{R} \Rightarrow$ pretpostavimo

$$X_j | Y=k \sim N(\mu_{k,j}, \sigma_{k,j}^2)$$

(kako da se... $\Sigma_k =$ diagonalna matrica...)

• ako $x_j \in S_j = \{0, \dots, K_j - 1\} \Rightarrow$ mpr, neparametarski

$$\hat{h}_{k_{ij}}(x) = \frac{\sum_{i=1}^n \mathbb{1}_{\{h_{k_{ij}} = k, x_{ij} = x\}}}{\sum_{i=1}^n \mathbb{1}_{\{h_{k_{ij}} = k\}}}$$

7.3.4 Usporedna metoda

• log. regresija — $\log\left(\frac{p_k(x)}{p_0(x)}\right) = \boxed{x^T \beta^{(k)}}$ (tipično $x_0 = 1$)

• LDA — $\log\left(\frac{p_k(x)}{p_0(x)}\right) \stackrel{DZ}{=} \log \frac{\pi_k}{\pi_0} = \frac{1}{2} (\mu_k + \mu_0)^T \Sigma^{-1} (\mu_k - \mu_0) + \boxed{x^T \Sigma^{-1} (\mu_k - \mu_0)}$

$=: \alpha_{k_0} + \sum_{i=1}^{K-1} \alpha_{k_i} \cdot x_i$

$\stackrel{x_0=1}{=} \boxed{x^T \alpha^{(k)}}$

↳ ipak, $\hat{\alpha}^{(m)} \neq \hat{\beta}^{(m)}$ [te je log. regresija općenitija jer ne pretpostavlja ništa nu rezd. $\mathbb{P}(X \in \cdot)$].

• NB — $\log\left(\frac{p_k(x)}{p_0(x)}\right) \stackrel{(7.15)}{=} \log\left(\frac{\pi_k}{\pi_0}\right) + \sum_{j=1}^P \log\left(\frac{h_{k_{ij}}(x_j)}{h_{0_{ij}}(x_j)}\right)$

$=: \alpha_{k_0} + \sum_{j=1}^P \alpha_{k_{ij}}(x_j)$

↳ flexibilne granice odlike

↳ ipak, redu interakcija

• QDA — $\log\left(\frac{p_k(x)}{p_0(x)}\right) = \alpha_k^1 + \sum_{j=1}^P \alpha_{k_{ij}} x_j + \boxed{\sum_{j_1, j_2=1}^P \alpha_{k_{ij_1 j_2}} x_{j_1} x_{j_2}}$

↑
interakcije

7.4 kNN metoda

↳ za $h_0 \in \{1, \dots, n\}$,

$$\hat{p}_h(x) = \frac{1}{h_0} \sum_{x^{(i)} \in H_{h_0}(x)} \mathbb{1}_{\{y_i = h\}}, \quad h \in S, x \in \mathbb{R}^p$$

$$f(x) = \underset{h \in S}{\operatorname{argmax}} \hat{p}_h(x) = \underset{h \in S}{\operatorname{argmax}} \sum_{x^{(i)} \in H_{h_0}(x)} \mathbb{1}_{\{y_i = h\}}$$

Uzpr. | za velike p

→ problem dimenzionalnosti!

"majority vote"
[uz 0-1 gubitak]

→ S-and-P.R