

7.2 Centralni gresioni teorem (CGT)

Koristimo da sl. var. X ima normalna (ili Gaussova)
razdistribuciju s parametrima μ i σ^2 ($\mu \in \mathbb{R}, \sigma > 0$),
ako je neprekidna s gustoćom

$$f_{N(\mu, \sigma^2)}(t) := \left[f(t) = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(t-\mu)^2}{2\sigma^2}}, t \in \mathbb{R} \right] \quad (7.6)$$

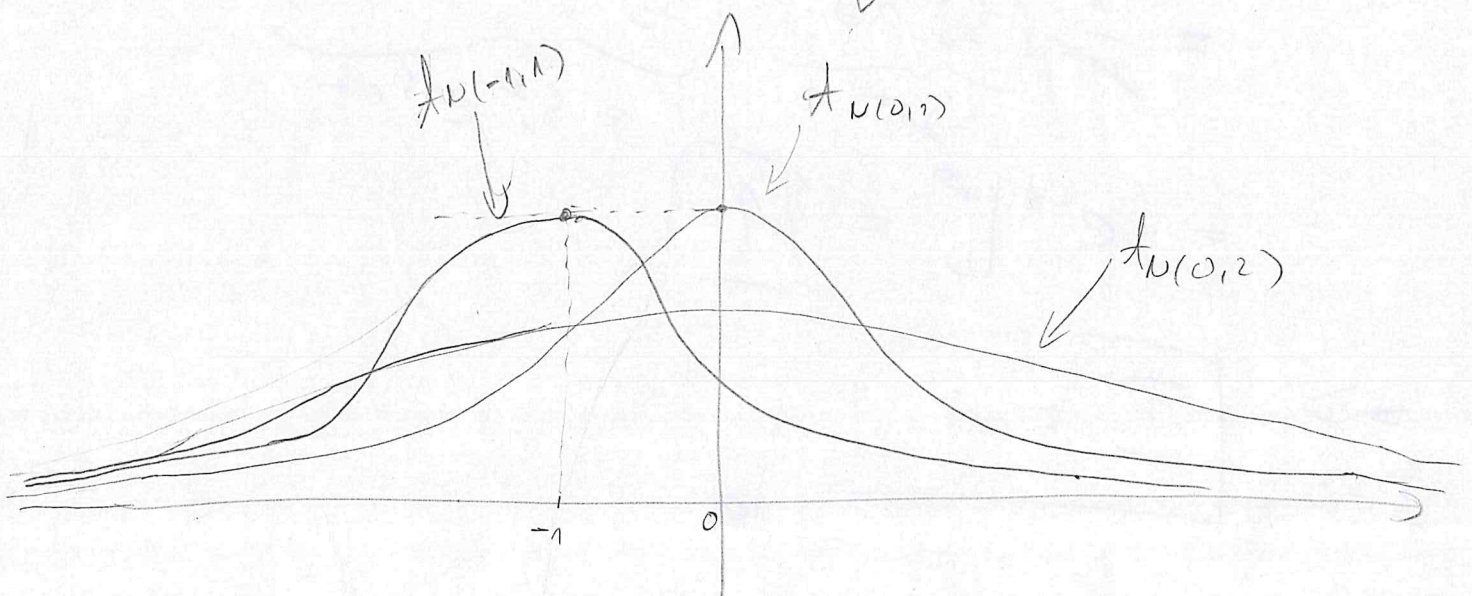
(oznaka $X \sim N(\mu, \sigma^2)$)

Razdistribuciju $N(0, 1)$ nazivamo standardna normalna
razdistribucija, a njenu f -ju distribucije obično

označavamo Φ

$$\Phi(x) = \int_{-\infty}^x \underbrace{\frac{1}{\sqrt{2\pi}} e^{-t^2/2}}_{= f_{N(0,1)}(t)} dt, x \in \mathbb{R} \quad (7.7)$$

[ne postoji za toveći deo!]



Nop. 7.31 Zauista unjedi

$$\underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{(t-\mu)^2}{2\sigma^2}} dt}_{=: I} = 1, \quad \forall \mu, \sigma^2.$$

Dokaz [02]

Imamo

$$I = \frac{1}{\sqrt{2\pi} \sigma} \int_{-\infty}^{\infty} e^{-\frac{(t-\mu)^2}{2\sigma^2}} dt$$

$$= \left[z = \frac{t-\mu}{\sigma}, \quad dz = \frac{dt}{\sigma} \right]$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-z^2/2} dz$$

$$\Rightarrow I^2 = \frac{1}{2\pi} \left(\int_{-\infty}^{\infty} e^{-z^2/2} dz \right) \cdot \left(\int_{-\infty}^{\infty} e^{-y^2/2} dy \right)$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(z^2+y^2)} dz dy$$

[polarne koordinate]

$$= \frac{1}{2\pi} \int_0^{2\pi} \int_0^{\infty} e^{-r^2/2} r dr d\theta$$

$$= \frac{1}{2\pi} \underbrace{\left(\int_0^{2\pi} d\theta \right)}_{=2\pi} \cdot \underbrace{\left(\int_0^{\infty} e^{-r^2/2} r dr \right)}_{= \int_0^{\infty} e^{-t} dt = -e^{-t} \Big|_0^{\infty} = 1}$$

$$= -e^{-t} \Big|_0^{\infty} = \boxed{1}$$

\Rightarrow $I \geq 0$ $I = 1$

□

Ako je $Z \sim N(0,1)$, imamo $\int_{-\infty}^{+\infty} \dots dt$ apsolutno integrabilno!

$$\bullet E[Z] = \int_{-\infty}^{+\infty} t \underbrace{\frac{1}{\sqrt{\pi}} e^{-t^2/2}}_{\text{neparna f-ja}} dt = \boxed{0} \quad (7.9)$$

$$\bullet \text{Var}(Z) = E[Z^2] = \int_{-\infty}^{+\infty} t^2 \frac{1}{\sqrt{\pi}} e^{-t^2/2} dt$$

$$= \left[\begin{array}{l} u = t \\ du = 1 \end{array} \quad \left. \begin{array}{l} dv = t e^{-t^2/2} \\ v = -e^{-t^2/2} \end{array} \right\} \right]$$

$$= \frac{-1}{\sqrt{\pi}} t e^{-t^2/2} \Big|_{-\infty}^{+\infty} + \frac{1}{\sqrt{\pi}} \int_{-\infty}^{+\infty} e^{-t^2/2} dt$$

$$= -0 + 0 + 1 = \boxed{1} \quad (7.10)$$

Prop. 7.10 (1) Ako je $Z \sim N(0,1)$, $\forall \mu \in \mathbb{R}$, $\sigma > 0$ vrijedi:

$$X := \sigma \cdot Z + \mu \sim N(\mu, \sigma^2) \quad (7.10)$$

Dokaz $\forall x \in \mathbb{R}$,

$$P(X \leq x) = P\left(Z \leq \frac{x-\mu}{\sigma}\right) \stackrel{Z \sim N(0,1)}{=} \int_{-\infty}^{\frac{x-\mu}{\sigma}} \frac{1}{\sqrt{\pi}} e^{-t^2/2} dt$$

$$= \int_{t=\frac{x-\mu}{\sigma}}^{z=\sigma \cdot t + \mu} \frac{1}{\sigma \sqrt{\pi}} e^{-\frac{(z-\mu)^2}{2\sigma^2}} dz = f_{N(\mu, \sigma^2)}(z)$$

\Rightarrow po def. $X \sim N(\mu, \sigma^2)$

(7.8) - (7.10) pravice da za $X \sim N(\mu, \sigma^2)$, vrijedi

$$\begin{aligned} E[X] &= \mu + \sigma \underbrace{E[Z]}_{=0} = \mu, \\ \text{Var}(X) &= \sigma^2 \cdot \underbrace{\text{Var}(Z)}_{=1} = \sigma^2, \end{aligned}$$

te

$$\frac{X - \mu}{\sigma} \sim N(0, 1)$$

Teor. 7.11 Neka su X_1, X_2, \dots njez. sl. varijablae t.d.

(CGT)

$$\sigma_i^2 = \text{Var}(X_i) < \infty,$$

te neka je $\mu_i = E[X_i]$, i $S_n = \sum_{i=1}^n X_i, \forall n \in \mathbb{N}$.

Tada vrijedi

$$(7.11) \quad \lim_{n \rightarrow \infty} P\left(\frac{S_n - n\mu}{\sqrt{n} \cdot \sigma} \leq x \right) = P(Z \leq x), \forall x \in \mathbb{R}$$

$\stackrel{= E[S_n]}{\text{}} \quad \sqrt{\text{Var}(S_n)} \quad (\text{bez dokaza})$

gdje je $Z \sim N(0, 1)$

"konvergencija po distribuciji"

Prop. 7.12 (7.11) pisemo

$$\frac{S_n - n\mu}{\sqrt{n} \sigma} \xrightarrow[n \rightarrow \infty]{d} Z \sim N(0, 1)$$

Uočimo, za "veliki" n i sve $x \in \mathbb{R}$,

$$P(S_n \leq x) = P\left(\frac{S_n - n\mu}{\sqrt{n}\sigma} \leq \frac{x - n\mu}{\sqrt{n}\sigma}\right)$$

$$\approx P\left(Z \leq \frac{x - n\mu}{\sqrt{n}\sigma}\right)$$

CGT, 15. (7.11)

$$= P\left(\underbrace{\sqrt{n}\sigma Z + n\mu}_{\sim N(n\mu, n\sigma^2)} \leq x\right),$$

tz. intuitivno,

$$S_n \stackrel{d}{\approx} N(n\mu, n\sigma^2)$$

"približno po distr."

" $E[S_n]$ $Var(S_n)$ "

Nije bitna distrib. od X_i , već samo da li je $\sigma^2 < \infty$

Prop. 7.13 Može se pokazati da (7.11) postaje

$$\lim_{n \rightarrow \infty} P\left(a \leq \frac{S_n - n\mu}{\sqrt{n}\sigma} \leq b\right) = P\left(a \leq Z \leq b\right) = \Phi(b) - \Phi(a),$$

(7.12)

za sve $-\infty < a < b < \infty$.

Pr. 7.14 Neka je $Y_n \sim B(n, p)$, $\forall n \in \mathbb{N}$, za $p \in (0, 1)$ fiksno.

Ako u Tm. 7.11 uzmemo $X_1 \sim \begin{pmatrix} 0 & 1 \\ p & p \end{pmatrix}$

$$\Rightarrow \mu = p, \sigma^2 = pq, \text{ te } S_n \sim Y_n.$$

$$\Rightarrow \frac{Y_n - np}{\sqrt{npq}} \xrightarrow[n \rightarrow \infty]{d} Z \sim N(0, 1) \quad (7.13)$$

(CGT)

ter. de Moivre-Laplaceov + m.

Moćima,

$$\frac{S_n - n\mu}{\sigma\sqrt{n}} \cdot \frac{1/n}{1/n} = \frac{\sqrt{n}}{\sigma} \left(\frac{S_n}{n} - \mu \right) \quad (7.14)$$

\Rightarrow CGT kaže koliko $\frac{S_n}{n}$ odstupa od μ

(po JZVB, $\frac{S_n}{n} \xrightarrow[n \rightarrow \infty]{g.o.} \mu$).

Pr. 7.15 | Nepravdi postotak glasača, p , glasat će za jednu stranku. Situaciju je odobrenu n ljudi te među je

$S_n = \#$ ispitanika koji će glasati za tu stranku.

Koliki treba biti n da $\frac{S_n}{n}$ bude na $\pm 3\%$ od p s vjerojatnošću od barem 0.95 ? [TV prijelaz]

(P) Uz pretpostavku da su ispitanici nezavisni,

$$S_n \sim B(n, p).$$

Želimo

$$0.95 = P\left(\left| \frac{S_n}{n} - p \right| \leq 0.03 \right)$$

$$= P\left(\frac{\sqrt{n}}{\sqrt{pq}} \left| \frac{S_n}{n} - p \right| \leq \frac{\sqrt{n}}{\sqrt{pq}} 0.03 \right)$$

$$\stackrel{\text{a.}}{\approx} \underbrace{P}_{(7.12) - (7.14)} \left(\Phi(t_n) - \Phi(-t_n) \right)$$

$$= 2\Phi(t_n) - 1,$$

$$\Phi(-t_n) = 1 - \Phi(t_n)$$

$$b. \quad \Phi(z_{\alpha}) \cong 0.975$$

$$\Leftrightarrow \Phi(1.96) \cong 0.975, \quad \text{imamo}$$

$$z_{\alpha} \cong 1.96,$$

me znamo $p!$

$$j. \quad \sqrt{n} \cong \frac{1.96}{0.03} \sqrt{pq} = 65.33 \sqrt{pq}. \quad (*)$$

Ipak, znamo $\sqrt{pq} = \sqrt{p(1-p)} \leq \sqrt{\frac{1}{2}(1-\frac{1}{2})} = \frac{1}{2}, \quad \forall p \in (0,1),$

pa onda je $n \cong 65.33^2 \cdot (\frac{1}{2})^2 = \boxed{1067.11},$

injeđi (*) pa onda:

$$P\left(\left|\frac{S_n}{n} - p\right| \leq 0.03\right) \geq 0.95$$

□

